# Free deconvolution- A tribute for Elisabeth work on deconvolution

Fabrice Gamboa

Work in progress with: Reda Chhalibi, Slim Kammoun and Mauricio Velasco Elisabeth Gassiat - a path in modern statistics

June 2023



One of Elisabeth's favourite themes : Deconvolution	Sparsity and discrete mixtures	Free multiplicative deconvolution	RMT for Large Cova

## Agenda



- Sparsity and discrete mixtures
- Free multiplicative deconvolution
- RMT for Large Covariance Matrices
  - Brute force SVD is wrong in large dimension for spectral recovery
  - Free deconvolution and the good transform
  - Our method

One of Elisabeth's	favourite themes	Deconvolution
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## Overview



Sparsity and discrete mixtures

Free multiplicative deconvolution

RMT for Large Covariance/Matrices

Sparsity and discrete mixtures Free multiplicative deconvolution

RMT for Large Cova

One of Elisabeth's favourite themes : Blind deconvolution

What is deconvolution?

 $X \sim P, Y \sim Q$  and  $Z = X + Y \sim R$ .

Knowing P and Q  $\Rightarrow$  compute R. Direct problem Convolution 

Knowing P and R  $\Rightarrow$  compute Q. Inverse problem Deconvolution 

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## Once upon the time, blind deconvolution

#### Déconvolution aveugle

par Elisabeth Gassiat



Description en francais

Description en anglais

#### Titre traduit

Blind deconvolution

#### Résumé

Considering a signal X which is a process of random variables identically independently distributed, and the signal Y obtained by filtering X through a linear system s, we study the estimation of s from the observation of y in the following semi-parametric situation the law of X is unknown and non Gaussian, and s has an inverse of convolution with finite length. We need no assumption on the phase of the system, i. E. On the causality or non causality of s. We propose an estimation by maximum objective. The estimates are consistent and asymptotically Gaussian this result is still available what-ever the dimension of the index space of the series is. We study the asymptotic efficiency of the estimate and, in the causal case, we compare it to the usual minimum square estimates. The output y being an autoregressive field, we propose a consis- tent method of identification of the order of the model. We study different types of robustness robustness to underparametrization, robustness to additive noise on the observations. We also inves tigate the case where the law of X has infinite moments, and we show that, for "standardized cumulants" as objectives, and under assumptions which are in particular verified for laws in the attraction demains of stable laws, the obtained estimates are still consistent, and the speed of convergence is, in the causal case, better than for laws with finite variance



## Elisabeth Ph. D main results



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## Elisabeth Ph. D main results

### Frame in a nutshell :

- $(Z_n)_{n \in \mathbb{Z}^d}$  observed A.R. built on an i.i.d. non Gaussian sequence  $(X_n)_{n \in \mathbb{Z}^d},$
- $b^* = (b_i^*)_{j \in S}$  the unknown weight vector  $(|S| = m < \infty)$ ,

• 
$$Y_n(b^*) := \sum_{j \in S} b_j^* Z_{n-j}$$
,

$$Z_n = Y_n(b^*) + X_n, \ n \in \mathbb{Z}^d$$

Observing only  $Z_1, \ldots, Z_n$  recover both  $b^*$  and  $X_1, \ldots, X_n$  $\Rightarrow$ Blind deconvolution (much more greedy than the deconvolution !!) One of Elisabeth's favourite themes : Deconvolution 000000<u>00</u>

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## Elisabeth Ph. D main results

Results in a nutshell :

- Build  $\hat{b}$  as an M-estimator of  $b^*$ ,
- The good M functions are the ones decreasing by convolution (typically normalized cumulants),
- Under very weak assumptions (even no variance for X!!) convergence and asymptotic rates for b

RMT for Large Cova

## One step beyond : optimality

The Annals of Statistics 1993, Vol. 21, No. 4, 2022-2042

### ADAPTIVE ESTIMATION IN NONCAUSAL STATIONARY AR PROCESSES

By E. GASSIAT

Université Paris-Sud

We consider the estimation problem of the parameter b of a stationary AR(p) process without any of the usual causality assumptions. The aim of the paper is to derive asymptotic minimax bounds for estimators of b. When the distribution of the noise is known, we show LAN properties of the model and derive locally asymptotically minimax (LAM) estimators. The most important results are about the case of unknown distribution: The main result shows that, if one uses the usual parametrization, these bounds depend heavily on the causality or the noncausality of the process, so that adaptive efficient estimation is impossible in the noncausal situation: The scaling factor is shown to give the hardest one-dimensional subproblem, and an unusual scaling is exhibited that could lead to adaptive efficient estimation of the rescaled parameter even in the noncausal case.

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## One step beyond : optimality

Results in a nutshell :

 For the estimation of b<sup>\*</sup> under assumptions on the density of X the model is LAN

$$\log\left(\frac{\mathsf{d}\mathsf{L}_{b^*+\frac{h}{\sqrt{n}}}}{\mathsf{d}\mathsf{L}_{b^*}}\right) = \langle \mathsf{h}, \Delta_{\mathsf{n}} \rangle - \frac{1}{2}\mathsf{h}^\mathsf{T}\mathsf{I}_{b^*}\mathsf{h} + \mathsf{o}_p(1),$$

with 
$$\Delta_n \stackrel{\mathcal{L}}{\Rightarrow} \mathcal{N}_m(0, I_{b^*})$$

- First consequence : lower bound for the asymptotic variance of a regular estimator (b is regular).
- $\hat{b}$  may be modified using one step Newton method to built asymptotic optimal b

One of Elisabeth's	favourite	themes	Deconvolution

## Overview

- One of Elisabeth's lavourite themes: Deconvolution
- Sparsity and discrete mixtures
- Free multiplicative deconvolution
- 4 RMT for Large Covariance Matrices

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## Uniqueness and stability in the generalized moment problem

C. R. Acad. Sci. Paris, t. 310, Série I, p. 41-44, 1990

Analyse mathématique/Mathematical Analysis

### Problèmes des moments et concentration de mesure

### Elisabeth GASSIAT

Résumé - Nous nous intéressons au problème des moments au voisinage des points singuliers. Nous mettons en évidence des hypothèses portant sur les fonctions de moments sous lesquelles : nous étudions tout d'abord l'information apportée par les moments sur la concentration de la mesure cherchée, et nous montrons que le diamètre de Prokhorov de l'ensemble des mesures qui sont solution du problème des moments tend vers 0 au voisinage des points singuliers. Ces hypothèses sont vérifiées lorsque les moments sont définis par des T-systèmes. Nous indiquons des applications en traitement du signal.

### The moment problem and the concentration of measures

Abstract – We study the moment problem for constraints in the vicinity of singular points. We indicate assumptions under which: we first study the information on the concentration of the measure given by its moments and we then show that the Prokhorov diameter of the convex set of the measures solution of the moment problem tends to 0, when the constraint tends to singular points. These assumptions hold for moments defined by T-systems. We give some ideas of applications in signal processing.

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## Uniqueness and stability in the Hausdorff moment problem

•  $v^*(dx) := \sum_{j=1}^{l} p_j^* \delta_{x_i^*}(dx) \in \mathcal{P}(]0, 1[)$  is completely determined by its 21 first algebraic moments

$$\mathfrak{m}_{j}(\mathbf{v}^{*}) = \int_{]0,1[} x^{j} \mathbf{v}^{*}(dx), \ j = 1 \dots 2l.$$

- $\bullet \ \text{Hankel Matrix}: H_{p+1}(\nu) = (\mathfrak{m}_{i+j-2}(\nu))_{1\leqslant i,j\leqslant p+1}, \nu\in \mathfrak{P}(]0,1[),$ 
  - $H_{n+1}(v) \ge 0$
  - det  $H_{p+1}(v) = 0 \Leftrightarrow v$  is supported by less than p+1 points.
  - So that det  $H_{p+1}$  can be used as M-function for deconvolution if the noise is discretely supported !!

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## Blind discrete deconvolution

The Annals of Statistics 1996, Vol. 24, No. 5, 1964-1981

#### BLIND DECONVOLUTION OF DISCRETE LINEAR SYSTEMS

By F. GAMBOA AND E. GASSIAT

### Laboratoire de Statistiques Orsav and Université Paris-Nord, and Université d'Evry Val d'Essonne

We study the blind deconvolution problem in the case where the input noise has a finite discrete support and the transfer linear system is not necessarily minimum phase. We propose a new family of estimators built using algebraic considerations. The estimates are consistent under very wide assumptions: The input signal need not be independently distributed: the cardinality of the finite support may be estimated simultaneously. We consider in particular AR systems: In this case, we prove that the estimator of the parameters is perfect a.s. with a finite number of observations.

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## Extension to source separation and mixtures



### Source Separation when the Input Sources Are Discrete or Have Constant Modulus

Fabrice Gamboa and Elisabeth Gassiat

Abstract - In this paper, we present a new method for the source separation problem when some prior information on the input sources is available. More specifically, we study the situation where the distributions of the input signals are discrete or are concentrated on a circle. The method is based on easy properties of Hankel forms and on the divisibility of Gaussian distributions. In both situations, we prove that the estimator converges in absence of noise or if we know the first moments of the noise up to its scale. Moreover, in the absence of noise, the estimate converges with a finite number of observations.

Index Terms - Estimation, Hankel forms, identification, independent sources, source separation.

The problem of interest is the restitution of the signal  $s(t), t = 1, \dots, N$  on the basis of the observations  $\mathbf{x}(t)$ .  $t = 1, \dots, N$ . This is the source separation problem. In absence of additive noise, it can be solved by estimating a decoupling matrix B to the observations, which may be seen as an estimator of some inverse matrix (up to scale and permutation) of A. Different separation procedures have been proposed in the past few years, for instance, using higher order statistics; see [2], [5], [6], [10], [15], [17], [22], [24], [25], [29], and [31]. Maximum likelihood estimates are considered in [1] and [11], whereas adaptive solutions are studied in [9], [18],

Bernoulli 3(3), 1997, 279-295

### The estimation of the order of a mixture model

#### DIDIER DACUNHA-CASTELLE<sup>1</sup> and ELISABETH GASSIAT<sup>2+</sup>

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<sup>2</sup>Equipe d'Analyse et de Probabilités, Université d'Evry, 91025 Evry Cédex, France

We propose a new method to estimate the number of different populations when a large sample of a mixture of these populations is observed. It is possible to define the number of different populations as the number of points in the support of the mixing distribution. For discrete distributions having a finite support, the number of support points can be characterized by Hankel matrices of the first algebraic moments, or Toeplitz matrices of the trigonometric moments. Namely, for one-dimensional distributions, the cardinality of the support may be proved to be the least integer such that the Hankel matrix (or the Toeplitz matrix) degenerates. Our estimator is based on this property. We first prove the convergence of the estimator, and then its exponential convergence under wide assumptions. The number of populations is not a priori bounded. Our method applies to a large number of models such as translation mixtures with known or unknown variance, scale mixtures, exponential families and various multivariate models. The method has an obvious computational advantage since it avoids any computation of estimates of the mixing parameters. Finally we give some numerical examples to illustrate the effectiveness of the method in the most popular cases.

Keywords: Hankel matrix: mixture models: order estimation: penalization

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RMT for Large Cova

## Recent contribution using this sparsity paradigm

The Annals of Statistics 2021 Vol 49 No 3 1779-1809 https://doi.org/10.1214/20-AOS2022 D Institute of Mathematical Statistics, 2021

### SUPERMIX: SPARSE REGULARIZATION FOR MIXTURES

BY Y. DE CASTRO<sup>1</sup>, S. GADAT<sup>2</sup>, C. MARTEAU<sup>3</sup> AND C. MAUGIS-RABUSSEAU<sup>4</sup>

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This paper investigates the statistical estimation of a discrete mixing measure  $\mu^0$  involved in a kernel mixture model. Using some recent advances in l1-regularization over the space of measures, we introduce a "data fitting and regularization" convex program for estimating  $\mu^0$  in a grid-less manner from a sample of mixture law, this method is referred to as Beurling-LASSO.

Our contribution is two-fold: we derive a lower bound on the bandwidth of our data fitting term depending only on the support of  $\mu^0$  and its socalled "minimum separation" to ensure quantitative support localization error bounds; and under a so-called "nondegenerate source condition" we derive a nonasymptotic support stability property. This latter shows that for a sufficiently large sample size n, our estimator has exactly as many weighted Dirac masses as the target  $\mu^0$ , converging in amplitude and localization towards the true ones. Finally, we also introduce some tractable algorithms for solving this convex program based on "Sliding Frank-Wolfe" or "Conic Particle Gradient Descent".

Statistical performances of this estimator are investigated designing a socalled "dual certificate", which is appropriate to our setting. Some classical situations as, for example, mixtures of super-smooth distributions (see, e.g., Gaussian distributions) or ordinary-smooth distributions (see, e.g., Laplace distributions), are discussed at the end of the paper.

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## Recent contribution of Elisabeth on deconvolution

The Annals of Statistics 2022, Vol. 50, No. 1, 303-323 https://doi.org/10.1214/21-AOS2106 D Institute of Mathematical Statistics, 2022.

### DECONVOLUTION WITH UNKNOWN NOISE DISTRIBUTION IS POSSIBLE FOR MULTIVARIATE SIGNALS

BY ÉLISABETH GASSIAT<sup>1</sup>, SYLVAIN LE CORFF<sup>2</sup> AND LUC LEHÉRICY<sup>3</sup>

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> This paper considers the deconvolution problem in the case where the target signal is multidimensional and no information is known about the noise distribution. More precisely, no assumption is made on the noise distribution and no samples are available to estimate it: the deconvolution problem is solved based only on observations of the corrupted signal. We establish the identifiability of the model up to translation when the signal has a Laplace transform with an exponential growth  $\rho$  smaller than 2 and when it can be decomposed into two dependent components. Then we propose an estimator of the probability density function of the signal, which is consistent for any unknown noise distribution with finite variance. We also prove rates of convergence and, as the estimator depends on p which is usually unknown, we propose a model selection procedure to obtain an adaptive estimator with the same rate of convergence as the estimator with a known tail parameter. This rate of convergence is known to be minimax when  $\rho = 1$ .

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Sparsity and discrete mixtures

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## Overview

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RMT for Large Covariance/Matrices

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## Classical convolutions

- X and Y real random variables,
- Z = X + Y classical additive convolution
- XY classical multiplicative convolution

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$$\mathbb{P}(X \in dx) := f(x) dx \quad \mathrm{and} \quad \mathbb{P}(Y \in dy) := g(y) dy$$

then

$$\mathbb{P}(Z \in dz) = f *_{\alpha} g(x) dx = \left( \int f(z - y)g(y) dy \right) dz \text{, and } \varphi_{Z} = \varphi_{X} \times \varphi_{Y}$$
$$\mathbb{P}(XY \in dz) = f *_{m} g(x) dx = \left( \int f(z/y)g(y) dy \right) dz \text{.}$$

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## Classical convolution with matrices

$$\begin{aligned} ^{(n)} \in M_n(\mathbb{C}) \\ \mu_{A^{(n)}} &:= \frac{1}{n} \sum_{\lambda \in \text{Spec}(A^{(n)})} \delta_{\lambda} \ \text{(spectral measure of } A^{(n)}) \end{aligned}$$

Two i.i.d. independent samples of positive random variables with laws  $\mu_1$ and  $\mu_2 X_1, X_2, X_3, \dots, X_n$  and  $Y_1, Y_2, Y_3, \dots, Y_n$ .  $A^{(n)} := \operatorname{diag}(X_1, \ldots, X_n)$  and  $B^{(n)} := \operatorname{diag}(Y_1, \ldots, Y_n)$ .

### Lemma (Obvious but illustrative)

We have

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$$\lim_{n \to \infty} \mu_{A^{(n)}} = \mu_1$$
 ,  $\qquad \lim_{n \to \infty} \mu_{B^{(n)}} = \mu_2$ 

and also

$$\lim_{n \to \infty} \mu_{A^{(n)}B^{(n)}} = \lim_{n \to \infty} \mu_{(B^{(n)})^{\frac{1}{2}}A^{(n)}(B^{(n)})^{\frac{1}{2}}} = \mu_1 *_m \mu_2.$$

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## Free multiplicative convolution

Two iid independent samples of positive random variables with laws  $\mu_1$ and  $\mu_2$ :

$$X_1, X_2, X_3, \ldots, X_n, \ldots \text{ and } Y_1, Y_2, Y_3, \ldots, Y_n, \ldots$$

$$A^{(n)} := \mathsf{diag}(X_1, \dots, X_n)$$
 and  $B^{(n)} := \mathbf{U_n} \operatorname{diag}(Y_1, \dots, Y_n) \mathbf{U_n^*}$ ,

 $U_n$  Haar on the orthogonal group (independent of X and Y).

### Theorem

Of course, we still have

$$\lim_{n\to\infty}\mu_{A^{(n)}}=\mu_1$$
 ,  $\qquad \lim_{n\to\infty}\mu_{B^{(n)}}=\mu_2$  ,

but this time :

$$\lim_{n \to \infty} \mu_{(B^{(n)})^{\frac{1}{2}} A^{(n)}(B^{(n)})^{\frac{1}{2}}} = \mu_1 \boxtimes \mu_2$$

Here  $\mu_1 \boxtimes \mu_2$  is a deterministic measure which depends only on  $\mu_1$  and  $\mu_2$  called their free multiplicative convolution.

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Free multiplicative deconvolution

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Sparsity and discrete mixtures Free multiplicative deconvolution

## Overview

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Free multiplicative deconvolution

RMT for Large Covariance Matrices

- Brute force SVD is wrong in large dimension for spectral recovery
- Free deconvolution and the good transform
- Our method

## Two pioneering papers

2008, Vol. 36, No. 6, 2757-2790 time of Mathematical Statistics, 200

2012, Vol. 40, No. 2, 1024-1060

#### NONLINEAR SHRINKAGE ESTIMATION OF LARGE-DIMENSIONAL COVARIANCE MATRICES

#### BY OLIVIER LEDOIT AND MICHAEL WOLF

#### University of Zurich

Many statistical applications require an estimate of a covariance matrix and/or its inverse. When the matrix dimension is large compared to the sample size, which happens frequently, the sample covariance matrix is known to perform poorly and may suffer from ill-conditioning. There already exists an extensive literature concerning improved estimators in such situations. In the absence of further knowledge about the structure of the true covariance matrix, the most successful approach so far, arguably, has been shrinkage estimation. Shrinking the sample covariance matrix to a multiple of the identity. by taking a weighted average of the two, turns out to be equivalent to linearly shrinking the sample eigenvalues to their grand mean, while retaining the sample eigenvectors. Our paper extends this approach by considering nonlinear transformations of the sample eigenvalues. We show how to construct an estimator that is asymptotically equivalent to an oracle estimator suggested in previous work. As demonstrated in extensive Monte Carlo simulations, the resulting bona fide estimator can result in sizeable improvements over the sample covariance matrix and also over linear shrinkage.

#### SPECTRUM ESTIMATION FOR LARGE DIMENSIONAL COVARIANCE MATRICES USING RANDOM MATRIX THEORY<sup>1</sup>

#### BY NOUREDDINE EL KAROUI

#### University of California, Berkeley

Estimating the eigenvalues of a population covariance matrix from a sample covariance matrix is a problem of fundamental importance in multivariate statistics; the eigenvalues of covariance matrices play a key role in many widely used techniques, in particular in principal component analysis (PCA). In many modern data analysis problems, statisticians are faced with large datasets where the sample size, n, is of the same order of magnitude as the number of variables p. Random matrix theory predicts that in this context, the eigenvalues of the sample covariance matrix are not good estimators of the eigenvalues of the population covariance.

We propose to use a fundamental result in random matrix theory, the Marčenko-Pastur equation, to better estimate the eigenvalues of large dimensional covariance matrices. The Marčenko-Pastur equation holds in very wide generality and under weak assumptions. The estimator we obtain can be thought of as "shrinking" in a nonlinear fashion the eigenvalues of the sample covariance matrix to estimate the population eigenvalues. Inspired by ideas of random matrix theory, we also suggest a change of point of view when thinking about estimation of high-dimensional vectors: we do not try to estimate directly the vectors but rather a probability measure that describes them. We think this is a theoretically more fruitful way to think about these problems.

Our estimator gives fast and good or very good results in extended simulations. Our algorithmic approach is based on convex optimization. We also show that the proposed estimator is consistent.

F. Gamboa (IMT Toulouse)

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## Our work in progress



### Mathematics > Probability

[Submitted on 9 May 2023]

### Estimation of large covariance matrices via free deconvolution: computational and statistical aspects

### Reda Chhaibi, Fabrice Gamboa, Slim Kammoun, Mauricio Velasco

The estimation of large covariance matrices has a high dimensional bias. Correcting for this bias can be reformulated via the tool of Free Probability Theory as a free deconvolution. The goal of this work is a computational and statistical resolution of this problem. Our approach is based on complex-analytic methods methods to invert S-transforms. In particular, one needs transforms live and an efficient computational scheme.

Comments: v1: Preliminary version Subjects: Probability (math.PR): Statistics Theory (math.ST): Computation (stat.CO) arXiv:2305.05646 [math.PR] Cite as: (or arXiv:2305.05646v1 [math.PR] for this version) https://doi.org/10.48550/arXiv.2305.05646

#### Submission history

From: Reda Chhaibi [view email] [v1] Tue, 9 May 2023 17:45:04 UTC (29 KB)

## Brute force SVD is wrong in large dimension

Consider a sequence of *centered* observables in  $\mathbb{R}^d$ , X<sub>1</sub>, X<sub>2</sub>,.... The empirical covariance

$$\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^{n} X_i X_i^* = \frac{1}{n} \mathbb{X} \mathbb{X}^* ,$$

where  $\mathbb{X}=(X_1,\ldots,X_n)\in M_{d,n}(\mathbb{R})$  is the matrix with columns given by the X<sub>i</sub>'s.

Take the  $X_i$  having iid components as well.

• If d fixed and  $n \to \infty$ .

$$\lim_{n o \infty} rac{1}{n} \mathbb{X} \mathbb{X}^* = \mathrm{I}_d$$
 ,

recovering the true covariance.

If both  $d = d_n$ ,  $n \to \infty$ , this is not true. •

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### Theorem (Marchenko-Pastur 1967)

Assume that  $d_n/n \to c < 1$ . Almost surely, as  $n \to \infty$ , we have the weak convergence of probability measures  $\lim_{n} \mu_{\perp_{XX^*}} = MP_c \neq \delta_1$  where

$$\mathsf{MP}_{\mathsf{c}}(\mathsf{d} x) := \mathbb{1}_{\{x \in [\mathfrak{l}, r]\}} \frac{\sqrt{(x-\mathfrak{l})(r-x)}}{2\pi x} \mathsf{d} x$$

is the Marchenko-Pastur distribution,  $l = (1 - \sqrt{c})^2$ ,  $r = (1 + \sqrt{c})^2$ .



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## The inverse problem

### Assume

$$\mathbb{Z} = \mathbb{X}\mathbb{Y} \text{ with } \mathbb{X} = \Sigma^{\frac{1}{2}},$$

with  $\Sigma \in M_d(\mathbb{R})$  true covariance, and  $\mathbb{Y} \in M_{d,n}(\mathbb{R})$  with iid entries with distribution  $\mathcal{N}(0, 1)$ .

### Question

When  $d/n \approx c > 0$ , how to estimate the spectrum of  $\Sigma$ ?

More precisely, we will focuss on the case where the limit spectral measure is V

$$\nu^* = \sum_{i=1}^n p_i^* \delta_{\sigma_i^*}$$

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## Model

Who is 
$$\nu^*$$
? 
$$\lim_{n \to +\infty} \frac{1}{d_n} \operatorname{Tr}(\Sigma^r) = \int x^r \nu(dx), \ (r \in \mathbb{N}).$$

Very roughly speaking,  $d_n p_i^*$  is the multiplicity of the eigenvalue  $\sigma_i^*$   $\Rightarrow$  Model not well defined

- Parametric model  $\theta^* = (\sigma_1^*, \dots, \sigma_k^*, p_1^*, \dots, p_{k-1}^*)^T$ ,
- As the white Gaussian is invariant under isometry take  $\Sigma$  diagonal,
- Diagonal elements  $\sigma_1^*, \ldots, \sigma_k^*$  with random multiplicity  $N_{1,n}, \ldots, N_{k,n}$   $(N_{1,n} + \cdots + N_{k,n} = d_n)$ ,

• 
$$\lim_{n\to\infty} \frac{N_{i,n}}{d_n} = p_i^*$$
 and  $n(\frac{N_{i,n}}{d_n} - p_i^*) = O_p(1).$ 

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## La La LAN

Under some assumptions on the random multiplicities :

### Proposition

The model is LAN for the rate  $n^{-1}$ 

$$\log\left(\frac{\mathsf{d} \mathsf{L}_{\theta^* + \frac{\mathsf{h}}{n}}}{\mathsf{d} \mathsf{L}_{\theta^*}}\right) = \langle \mathsf{h}, \Delta_{\mathsf{n}} \rangle - \frac{1}{2} \mathsf{h}^\mathsf{T} \mathsf{I}_{\theta^*} \mathsf{h} + \mathsf{o}_{\mathsf{p}}(1)$$

with 
$$\Delta_n \stackrel{\mathcal{L}}{\Rightarrow} \mathcal{N}_{2k-1}(0, I_{\theta^*})$$

Consequently, possible asymptotic efficient estimation could be possible (Newton Method). But

- $I_{\theta^*}$  is not yet completely computable !!
- Our estimate (see next slides) have to be regular (not shown yet) !!

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## Definitions and notations

• The Cauchy-Stieltjes transform of  $\mu$ , a probability measure on  $\mathbb{R}_+$  is :

 $\begin{array}{ccc} G_{\mu} \colon \mathbb{C}_{+} & \rightarrow & \mathbb{C}_{-} \\ z & \mapsto & \int_{\mathbb{R}_{+}} \frac{\mu(d\nu)}{z - \nu} \end{array}, \end{array}$ 

where

$$\mathbb{C}_{\pm} := \{z \in \mathbb{C} \mid \pm \Im z > 0\}$$
.

- The moment generating function is  $M_{\mu}(z) = zG_{\mu}(z) 1$ ٠
- For  $\mu \neq \delta_0$ ,  $M_{\mu}$  is invertible in the neighborhood of  $\infty$  and the inverse is denoted by  $M_{\mu}^{\langle -1 \rangle}$ . The S-transform of  $\mu$  is defined as

$$S_{\mu}(m) = rac{1+m}{mM_{\mu}^{\langle -1 \rangle}(m)},$$

and is analytic in a neighborhood of m = 0. (WARNING : Inverse is multi-valued).

• For a diagonalizable matrix  $A^{(N)} \in M_N(\mathbb{R})$ , we write  $S_{A^{(N)}} := S_{\mu_{A^{(N)}}}$  $G_{A^{(N)}} := G_{\mu_{A^{(N)}}}, M_{A^{(N)}} := M_{\mu_{A^{(N)}}}.$ 

## S-transforms and free convolution

### A more complete statement on free multiplicative convolution :

### Theorem (Voiculescu, 1987)

Consider two sequences of positive matrices, each element in  $M_{N}(\mathbb{R})$ 

$$\left(A^{\,(\,N\,\,)}\ ;\ N\geqslant 1\right)$$
 ,  $\left(B^{\,(\,N\,\,)}\ ;\ N\geqslant 1\right)$  ,

such that :

$$\lim_{N \to \infty} \mu_A(N) = \mu_A , \quad \lim_{N \to \infty} \mu_B(N) = \mu_B .$$

Under the (technical) assumption of asymptotic freeness for A (N) and B (N), there exists a deterministic probability measure  $\mu_A \boxtimes \mu_B$  such that :

$$\lim_{N \to \infty} \mu_{(A(N))\frac{1}{2}B(N)(A(N))\frac{1}{2}} = \mu_A \boxtimes \mu_B.$$

The operation is the multiplicative free convolution. Moreover

$$S_{\mu_{A}\boxtimes\mu_{B}}(\mathfrak{m}) = S_{\mu_{A}}(\mathfrak{m})S_{\mu_{B}}(\mathfrak{m}).$$
<sup>(1)</sup>

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## Recall $\mathbb{X} = \Sigma^{\frac{1}{2}} \mathbb{Y}$ so that :

$$\frac{1}{n} \mathbb{X} \mathbb{X}^* = \Sigma^{\frac{1}{2}} \left( \frac{1}{n} \mathbb{Y} \mathbb{Y}^* \right) \Sigma^{\frac{1}{2}} \ .$$

### Corollary

Assume that  $\mathbb{Y}$  has iid coefficients, and  $\Sigma$  has a spectral measure converging to  $\nu.$  Under our assumptions,  $\mu_{\frac{1}{n}\mathbb{X}\mathbb{X}^*}$  converges weakly to a measure u satisfying :

$$\mu = \nu \boxtimes MP_c$$
 .

 $\rightsquigarrow$  We want to estimate  $\nu$  by free deconvolution i.e. such that

$$S_{\nu}(m) = S_{\mu}(m)/S_{MP_c}(m)$$

In particular we want to construct an empirical version  $\hat{v}_n$  such that

$$S_{\hat{v}_n}(\mathfrak{m}) \approx S_{\frac{1}{n}\mathbb{XX}^*}(\mathfrak{m})/S_{MP_c}(\mathfrak{m})$$

Ingredients for the estimator  $\hat{\nu}_n$  :

### Proposition (Inversion)

It is possible to recover moments of  $\hat{v}_n$  via (numerically constructible and small) contour integrals implicating  $G_{\frac{1}{n}\mathbb{X}\mathbb{X}^*}$ .

### Proposition (Reconstruction)

From the observation of (noisy) moments  $m_1, m_2, \ldots, m_k$ , one can reconstruct an atomic probability measure  $\hat{\nu}_n$  having these moments.

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