

## Minimax Lower Bounds

### ① Minimax risk

- Statistical setting:

- $(P_f)_{f \in \mathcal{F}}$ : a set of distributions on a measurable space  $(Y, \mathcal{A})$
- $d$ : a distance on  $\mathcal{F}$
- risk: for any estimator  $\hat{f}: Y \rightarrow \mathcal{F}$ , we consider the risk  $R(\hat{f}) = \mathbb{E}_{\hat{f}} [d(\hat{f}(Y), f)^q]$  for some  $q > 0$ .

Ex:  $\mathcal{F} = \mathbb{R}^n$ ,  $P_f = N(f, \sigma^2 I_n)$ ,  $d = \text{Euclidean distance}$ ,  $q=2$

$$\left\{ R(\hat{f}) = \mathbb{E}_{\hat{f}} [\|\hat{f}(Y) - f\|^2] \right.$$

- Best estimator  $\hat{f}$ ?



For all  $f \in \mathcal{F}$ , we have

$$\min_{\hat{f}: Y \xrightarrow{\text{meas.}} \mathcal{F}} \mathbb{E}_{\hat{f}} [d(\hat{f}(Y), f)^q] = 0 \quad (\text{reached for } \hat{f}(Y) = f)$$

$\rightsquigarrow$  no sense



we want  $\hat{f}$  to be good on the whole class  $\mathcal{F}$

- Minimax risk:

$$R^*(\mathcal{F}) := \min_{\hat{f}: Y \xrightarrow{\text{meas.}} \mathcal{F}} \max_{f \in \mathcal{F}} \mathbb{E}_{\hat{f}} [d(\hat{f}(Y), f)^q]$$

- our goal: proving some lower bounds on  $R^*(\mathcal{F})$  -

- Useful?: If

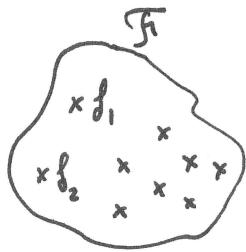
→ we prove  $R^*(\mathcal{F}) \geq \text{lower bound}$

→ and find  $\hat{f}$  such that  $R(\hat{f}) \approx \text{lower bound}$

then,  $\hat{f}$  performs almost as well as the best estimator in terms of minimax risk.

- Recipe:

- discretization of  $\mathcal{F}$ :  $\max_{f \in \mathcal{F}} \geq \max_{f \in \{f_1, \dots, f_N\}}$



- use lower bounds from information theory.

## ② A recipe for proving lower bounds (in 3 steps)

Step 1: a key lemma from information theory.

Kullback - Leibler divergence: For any  $P \ll Q$ , then

$$KL(P, Q) := \int \log \frac{dP}{dQ} dP \geq 0$$

Ex: Gaussian distribution  $P_f = \mathcal{N}(f, \sigma^2 I_m)$

$$\begin{aligned} KL(P_f, P_g) &= \int_{x \in \mathbb{R}^m} \log \frac{e^{-\|x-f\|^2/2\sigma^2}}{e^{-\|x-g\|^2/2\sigma^2}} dP_f(x) \\ &= \int_{x \in \mathbb{R}^m} \frac{1}{2\sigma^2} (\|f-g\|^2 + 2 \langle x-f, f-g \rangle) dP_f(x) \\ &= \frac{\|f-g\|^2}{2\sigma^2} \end{aligned}$$

Fano's Lemma

For any  $\bar{P}_1, \dots, \bar{P}_N, Q$  probability distribution on  $Y$ , such that  $\bar{P}_j \ll Q$ , for  $j=1, \dots, N$ , we have

$$\min_{\hat{\pi}: Y \rightarrow \{1, \dots, N\}} \frac{1}{N} \sum_{j=1}^N \bar{P}_j [\hat{\pi}(Y) \neq j] \geq 1 - \frac{1 + \frac{1}{N} \sum_{j=1}^N KL(\bar{P}_j, Q)}{\log(N)}$$

Remark: a classical choice for  $Q$  is  $Q = \frac{1}{N} \sum_{j=1}^N \bar{P}_j$

Proof:

- We first observe that

$$\min_{\hat{\pi}: Y \rightarrow \{1, \dots, N\}} \frac{1}{N} \sum_{j=1}^N \bar{P}_j [\hat{\pi}(Y) \neq j] = 1 - \max_{\hat{\pi}: Y \rightarrow \{1, \dots, N\}} \underbrace{\frac{1}{N} \sum_{j=1}^N \bar{P}_j [\hat{\pi}(Y) = j]}_{\text{to be upper-bounded}}$$

Lemma: explicit formula

$$\max_{\hat{\pi}: Y \rightarrow \{1, \dots, N\}} \frac{1}{N} \sum_{j=1}^N \bar{P}_j [\hat{\pi}(Y) = j] = \frac{1}{N} \mathbb{E}_Q \left[ \max_{j=1, \dots, N} \frac{d\bar{P}_j}{dQ}(Y) \right]$$

## Proof of formula:

$$\begin{aligned}
 \cdot \sum_{j=1}^N \hat{P}_j [\hat{\mathbb{I}}(Y)=j] &= \sum_{j=1}^N \int_Y \mathbb{I}_{\hat{\mathbb{I}}(y)=j} \underbrace{\frac{d\hat{P}_j}{dQ}(y)}_{\leq \max_{k=1,\dots,N} \frac{d\hat{P}_k}{dQ}(y)} dQ(y) \\
 &\leq \int_Y \underbrace{\sum_{j=1}^N \mathbb{I}_{\hat{\mathbb{I}}(y)=j}}_{=1} \times \max_{k=1,\dots,N} \frac{d\hat{P}_k}{dQ}(y) dQ(y) \\
 &= \mathbb{E}_Q \left[ \max_{k=1,\dots,N} \frac{d\hat{P}_k}{dQ}(Y) \right]
 \end{aligned}$$

- In addition, the inequality is an equality for the MLE

$$\hat{\mathbb{I}}(y) \in \operatorname{argmax}_{k=1,\dots,N} \frac{d\hat{P}_k}{dQ}(y)$$

□

- We can upper bound  $\mathbb{E}_Q \left[ \max_{j=1,\dots,N} \frac{d\hat{P}_j}{dQ}(Y) \right]$  with a lemma from Lecture 1

## Lemma (Lecture 1)

For any  $Z_1, \dots, Z_N$  random variables with value in an interval  $I \subset \mathbb{R}$ , and any  $\varphi: I \rightarrow \mathbb{R}^+$  convex, we have

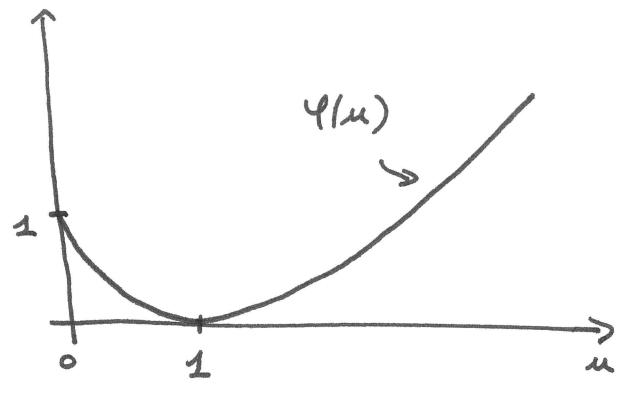
$$\varphi(\mathbb{E}[\max_{j=1,\dots,N} Z_j]) \leq \sum_{j=1}^N \mathbb{E}[\varphi(Z_j)]$$

• We choose

$$\varphi(u) = u \log u - u + 1$$

for  $u > 0$ .

$$\text{and } Z_j = \frac{d\bar{P}_j}{dQ}(y)$$



$$\begin{aligned} \cdot \mathbb{E}_Q \left[ \varphi \left( \frac{d\bar{P}_j}{dQ}(y) \right) \right] &= \int_y \log \left( \frac{d\bar{P}_j}{dQ}(y) \right) \underbrace{\frac{d\bar{P}_j}{dQ}(y) dQ(y)}_{d\bar{P}_j(y)} - \underbrace{\int_y \frac{d\bar{P}_j}{dQ}(y) dQ(y) + 1}_{= 1} \\ &= KL(\bar{P}_j, Q) \end{aligned}$$

so

$$\varphi \left( \mathbb{E}_Q \left[ \max_{j=1,\dots,N} \frac{d\bar{P}_j}{dQ}(y) \right] \right) \leq \sum_{j=1}^N KL(\bar{P}_j, Q)$$

$\underbrace{\phantom{\max_{j=1,\dots,N}}}_{=: Nu}$

$$\begin{aligned} \cdot \varphi(Nu) &= Nu(\log N + \log(u)) - Nu + 1 \\ &= Nu \log N + N \underbrace{(u \log u - u + 1)}_{\varphi(u) \geq 0} - (N-1) \\ &\geq Nu \log N - N \end{aligned}$$

So replacing  $u$  by its value :

$$\log(N) \times \mathbb{E}_Q \left[ \max_{j=1,\dots,N} \frac{d\bar{P}_j}{dQ}(y) \right] \leq N + \sum_{j=1}^N KL(\bar{P}_j, Q).$$

Conclusion:

$$\begin{aligned} \min_{\hat{\pi}: Y \rightarrow \{1,\dots,N\}} \frac{1}{N} \sum_{j=1}^N \bar{P}_j [\hat{\pi}(Y) \neq j] &= 1 - \frac{1}{N} \mathbb{E}_Q \left[ \max_{j=1,\dots,N} \frac{d\bar{P}_j}{dQ}(y) \right] \\ &\geq 1 - \frac{1}{\log(N)} \left( 1 + \frac{1}{N} \sum_{j=1}^N KL(\bar{P}_j, Q) \right) \end{aligned}$$

□

Step 2: From Fano's lemma to a lower bound over a finite set  $\{f_1, \dots, f_N\}$

- For any  $\hat{f}: Y \rightarrow \mathcal{F}_1$  measurable, we define

$$\hat{\pi}(y) \in \operatorname{argmin}_{j=1, \dots, N} d(f_j(y), f_j)$$

- We have  $\forall j$ :

$$\begin{aligned} \min_{i \neq k} d(f_i, f_k) \cdot \mathbb{P}_{\hat{\pi}(y) \neq j} &\leq d(f_j, f_{\hat{\pi}(y)}) \\ &\leq d(f_j, \hat{f}(y)) + d(\hat{f}(y), f_{\hat{\pi}(y)}) \\ \text{definition of } \hat{\pi} &\leq 2 d(f_j, \hat{f}(y)) - \end{aligned}$$

So, for any  $\hat{f}$ :

$$\begin{aligned} \frac{1}{N} \sum_{j=1}^N \mathbb{E}_{f_j} [d(f_j, \hat{f}(Y))^q] &\geq \frac{1}{2^q} \min_{i \neq k} d(f_i, f_k)^q \times \frac{1}{N} \sum_{j=1}^N \mathbb{P}_{f_j} [\hat{\pi}(Y) \neq j] \\ &\geq \frac{1}{2^q} \min_{i \neq k} d(f_i, f_k)^q \times \min_{\hat{\pi}: Y \rightarrow \{1, \dots, N\}} \frac{1}{N} \sum_{j=1}^N \mathbb{P}_{f_j} [\hat{\pi}(Y) \neq j] \end{aligned}$$

Hence, the corollary from Fano's lemma

### Corollary 3.4

For any  $\{f_1, \dots, f_N\} \subset \mathcal{F}_1$  and  $\mathcal{Q} \gg \mathbb{P}_{f_j}, j=1, \dots, N$

$$\min_{\hat{f}: Y \rightarrow \mathcal{F}_1} \frac{1}{N} \sum_{j=1}^N \mathbb{E}_{f_j} [d(f_j, \hat{f}(Y))^q]$$

$$\geq \frac{1}{2^q} \min_{i \neq k} d(f_i, f_k)^q \times \left( 1 - \frac{1 + \frac{1}{N} \sum_{j=1}^N \text{KL}(\mathbb{P}_{f_j}, \mathcal{Q})}{\log(N)} \right)$$

Step 3: finding a good discretization

For any  $\{f_1, \dots, f_N\} \subset \mathcal{F}$ :

$$\begin{aligned}
 R^*(\mathcal{F}) &:= \min_{\hat{f}: Y \rightarrow \mathcal{F}} \max_{f \in \mathcal{F}} \mathbb{E}_f [d(\hat{f}(Y), f)^q] \\
 &\geq \min_{\hat{f}: Y \rightarrow \mathcal{F}} \max_{j=1, \dots, N} \mathbb{E}_{f_j} [d(\hat{f}(Y), f_j)^q] \\
 &\geq \min_{\hat{f}: Y \rightarrow \mathcal{F}} \frac{1}{N} \sum_{j=1}^N \mathbb{E}_{f_j} [d(\hat{f}(Y), f_j)^q] \\
 &\geq \frac{1}{2^q} \min_{i \neq k} d(f_i, f_k)^q \times \left(1 - \frac{1 + \frac{1}{N} \sum_{j=1}^N KL(P_{f_j}, Q)}{\log(N)}\right) \quad (\text{Cor. 3.4})
 \end{aligned}$$

All the art is to find a good discretization  $\{f_1, \dots, f_N\}$ .

balance between

$$\begin{aligned}
 &\min_{i \neq k} d(f_i, f_k) \text{ as large as possible} \\
 &\frac{1 + \frac{1}{N} \sum_{j=1}^N KL(P_{f_j}, Q)}{\log(N)} \text{ smaller than 1}
 \end{aligned}$$

recipe: find  $f_1, \dots, f_N$  with

$$\left\{ \frac{1 + \frac{1}{N} \sum_{j=1}^N KL(P_{f_j}, Q)}{\log(N)} \leq \frac{1}{2} \quad \text{and} \quad d(f_i, f_k) \text{ as large as possible} \right.$$

Remark: there is a variant of Fano's lemma (based on Binge's Lemma) which is sometimes more handy - It leads to the next variant of Corollary 3.4

Corollary 3.6:

For any  $\{f_1, \dots, f_N\} \subset \mathcal{F}$  such that

$$\max_{j \neq k} KL(\mathbb{P}_{f_j}, \mathbb{P}_{f_k}) \leq \frac{2e}{2e+1} \log(N) \quad (*)$$

We have

$$\min_{\hat{f}: Y \rightarrow \mathcal{F}} \max_{j=1, \dots, N} \mathbb{E}_{f_j} [d(\hat{f}(Y), f_j)^q] \geq \frac{1}{2^q(2e+1)} \min_{j \neq k} d(f_j, f_k)^q.$$

Proof: With the notations of Fano's lemma, the events  $A_j = \{\hat{\mathcal{I}}(Y) = j\}$  are disjoint so:

Theorem B.13 ensures that

$$\min_{j=1, \dots, N} \mathbb{P}_j [\hat{\mathcal{I}}(Y) = j] \leq \frac{2e}{2e+1} \vee \max_{j \neq k} \frac{KL(\mathbb{P}_{f_j}, \mathbb{P}_{f_k})}{\log(N)}$$

$$(*) \rightarrow \leq \frac{2e}{2e+1}$$

Hence we get the variant of Fano's lemma: when  $(*)$  holds

$$\min_{\hat{\mathcal{I}}: Y \rightarrow \{1, \dots, N\}} \max_{j=1, \dots, N} \mathbb{P}_{f_j} [\hat{\mathcal{I}}(Y) \neq j] \geq \frac{1}{2e+1} -$$

Conclusion: same lines as proof of Corollary 3.4. □

### ③ Minimax risk for coordinate sparse regression

- We consider here  $\mathcal{F}_D = \{f = X\beta : \|\beta\|_0 \leq D\}$ ,
- $P_f = \mathcal{N}(f, \sigma^2 I_m)$ ,  $d(f_1, f_2) = \|f_1 - f_2\|$  and  $q=2$ .
- we have seen that  $KL(P_{f_1}, P_{f_2}) = \frac{\|f_1 - f_2\|^2}{2\sigma^2}$
- Restricted isometry constants: for  $D_{\max} \leq p/2$

$$\underline{c}_x := \inf_{\|\beta\|_0 \leq 2D_{\max}} \frac{\|X\beta\|}{\|\beta\|} \leq \sup_{\|\beta\|_0 \leq 2D_{\max}} \frac{\|X\beta\|}{\|\beta\|} =: \bar{c}_x$$

#### Theorem 3.5

- Let us fix  $D_{\max} \leq p/5$ .
- For any  $D \leq D_{\max}$ , we have

$$R^*(\mathcal{F}_D) \geq \frac{e}{4(2e+1)^2} \left(\frac{\underline{c}_x}{\bar{c}_x}\right)^2 \times D \log\left(\frac{p}{5D}\right) \times \sigma^2$$

Proof: The recipe is to

- find  $f_1, \dots, f_N \in \mathcal{F}_D$ , well spread and fulfilling (x)
- apply corollary 3.6

Lemma 3.7:

For any  $D \leq p/s$ , there exists  $\{\beta_1, \dots, \beta_N\} \subset \{\beta \in \mathbb{R}^p : \|\beta\|_0 = D\}$  such that

$$\text{i)} \quad \|\beta_j - \beta_k\|_0 \geq D, \quad \forall j \neq k$$

$$\text{ii)} \quad \log N \geq \frac{D}{2} \log \frac{p}{5D}$$

proof: exercise 3.6.2 □

- We choose  $f_j = r \times \beta_j$ ,  $j=1, \dots, N$ , with  $r$  such that (\*) holds:

$$\begin{aligned} \max_{j \neq k} KL(\mathbb{P}_{f_j}, \mathbb{P}_{f_k}) &= \max_{j \neq k} \frac{r^2 \|X(\beta_j - \beta_k)\|^2}{2\sigma^2} \\ &\leq \frac{r^2}{2\sigma^2} \bar{C}_X^2 \max_{j \neq k} \underbrace{\|\beta_j - \beta_k\|_0^2}_{\stackrel{\text{scaling}}{\downarrow}} \\ &\leq \|\beta_j - \beta_k\|_0 \leq 2D \\ &\stackrel{(*)}{\leq} \frac{2e}{2e+1} \log N \end{aligned}$$

$$\text{OK for } r^2 = \frac{\sigma^2}{\bar{C}_X^2 D} \times \frac{2e}{2e+1} \log N$$

- In addition:

$$\begin{aligned} \|f_j - f_k\|^2 &= r^2 \|X(\beta_j - \beta_k)\|^2 \\ &\geq r^2 \bar{C}_X^2 \underbrace{\|\beta_j - \beta_k\|_0^2}_{\stackrel{\text{ii)}}{\geq D}} &\geq r^2 D \bar{C}_X^2 \\ &= \|\beta_j - \beta_k\|_0 \stackrel{\text{ii)}}{\geq D} \end{aligned}$$

$$\geq \left(\frac{\bar{C}_X}{\bar{C}_X}\right)^2 \sigma^2 \times \frac{2e}{2e+1} \log N \geq \left(\frac{\bar{C}_X}{\bar{C}_X}\right)^2 \sigma^2 \frac{e}{2e+1} \times D \log \frac{p}{5D}$$

- Applying Corollary 3.6 gives Theorem 3.5 □