

## HOMEWORK: CONVERGENCE OF RANDOM VARIABLES AND CHARACTERISTIC FUNCTIONS

### 1. LÉVY'S CRITERION OF CONVERGENCE IN LAW

In this part,  $(X_n)_{n \in \mathbb{N}}$  is a sequence of real random variables; their laws in  $\mathcal{M}^1(\mathbb{R})$  are denoted  $\mu_n$ , and their characteristic functions are denoted  $\phi_n(t) = \mu_n(e^{itx}) = \mathbb{E}[e^{itX_n}]$ .

- (1) Show that for any real valued random variable  $X$ , the characteristic function  $\phi(t) = \mathbb{E}[e^{itX}]$  is a continuous function  $\mathbb{R} \rightarrow \mathbb{C}$ .

In the following we suppose that  $(\phi_n)_{n \in \mathbb{N}}$  converges pointwise to a function  $\phi : \mathbb{R} \rightarrow \mathbb{C}$  which is continuous at  $t = 0$ :

$$\forall t \in \mathbb{R}, \quad \lim_{n \rightarrow \infty} \phi_n(t) = \phi(t); \quad (1)$$

$$\text{and} \quad \lim_{t \rightarrow 0} \phi(t) = \phi(0) = 1. \quad (2)$$

Questions 2. to 4. deal with the tightness of the sequence of probability measures  $(\mu_n)_{n \in \mathbb{N}}$  under these hypotheses; Questions 5. to 8. deal with the unicity of a limit of a subsequence.

- (2) Using Fubini's theorem, show that for any law  $\mu \in \mathcal{M}^1(\mathbb{R})$  of characteristic function  $\phi$ , and any  $\varepsilon > 0$ ,

$$I_\varepsilon = \int_{-\varepsilon}^{\varepsilon} (1 - \phi(t)) dt = 2\varepsilon \int_{\mathbb{R}} \left(1 - \frac{\sin(\varepsilon x)}{\varepsilon x}\right) \mu(dx).$$

Show that if  $|\varepsilon x| \geq 2$ , then  $1 - \frac{\sin(\varepsilon x)}{\varepsilon x} \geq \frac{1}{2}$ . Deduce from it the inequality

$$\mu(\{x, |\varepsilon x| \geq 2\}) \leq \frac{I_\varepsilon}{\varepsilon}.$$

- (3) Fix  $\eta > 0$ . Under the hypotheses (1) and (2), show that there is an  $\varepsilon > 0$  and an integer  $N$  such that

$$\forall n \geq N, \quad \frac{I_{\varepsilon, n}}{\varepsilon} \leq \eta.$$

Conclude that (1) and (2) imply the tightness of the sequence of laws  $(\mu_n)_{n \in \mathbb{N}}$ .

- (4) Show that  $\phi$  is the characteristic function of a law  $\mu \in \mathcal{M}^1(\mathbb{R})$  (hint: use a convergent subsequence of  $(\mu_n)_{n \in \mathbb{N}}$ ).

- (5) Let  $\mu$  and  $\nu$  be two probability measures with characteristic functions  $\phi_\mu$  and  $\phi_\nu$ . Prove the Parseval identity

$$\int_{\mathbb{R}} e^{-itx} \phi_\mu(x) \nu(dx) = \int_{\mathbb{R}} \phi_\nu(y - t) \mu(dy). \quad (3)$$

- (6) Take  $\nu = \mathcal{N}(0, \varepsilon)$ , a Gaussian law of variance  $\varepsilon$  with density and characteristic function

$$\nu(dx) = \frac{1}{\sqrt{2\pi\varepsilon}} e^{-\frac{x^2}{2\varepsilon}} dx \quad ; \quad \phi_\nu(t) = e^{-\frac{\varepsilon t^2}{2}}.$$

Denote  $X$  a random variable under the law  $\mu$ , and  $Y_{\varepsilon^{-1}}$  an independent random variable under the law  $\mathcal{N}(0, \varepsilon^{-1})$ . Prove that the quantity of Equation (3), viewed as a function  $D(\varepsilon, t)$  of  $\varepsilon$  and  $t$ , is proportional to the density (in  $t$ ) of the law of  $X + Y_{\varepsilon^{-1}}$ .

- (7) Verify that  $Y_{\varepsilon^{-1}}$  converges in probability to the constant 0 as  $\varepsilon$  goes to infinity, and that  $X + Y_{\varepsilon^{-1}} \rightarrow X$ .
- (8) Show that if  $\phi_\mu$  is known, then so is the law  $\mu$ , so that  $\mu \mapsto \phi_\mu$  is injective.
- (9) Conclude that under the hypotheses (1) and (2), the laws  $\mu_n$  converge to a law  $\mu$  of a random variable (for the topology of convergence in law in  $\mathcal{M}^1(\mathbb{R})$ ). This is Lévy's continuity theorem.
- (10) Show the converse implication: if  $\mu_n \rightarrow \mu$ , then (1) and (2) hold with  $\phi_\mu = \phi$ .
- (11) Application: let  $B_n$  be a sequence of independent random Bernoulli variables with  $\mathbb{P}[B_n = 1] = \frac{1}{n}$  and  $\mathbb{P}[B_n = 0] = 1 - \frac{1}{n}$ . Recall that  $\sum_{k=1}^n \frac{1}{k} = \log n + O(1)$ . We set

$$\tilde{X}_n = \frac{\sum_{k=1}^n (B_k - \frac{1}{k})}{\sqrt{\log n}} \quad ; \quad X_n = \frac{(\sum_{k=1}^n B_k) - \log n}{\sqrt{\log n}}.$$

Show that

$$\mathbb{E}[e^{it\tilde{X}_n}] = \prod_{k=1}^n \left( 1 - \frac{t^2}{2k \log n} + O\left( \frac{t^2}{k^2 \log n} + \frac{t^3}{k (\log n)^{3/2}} \right) \right),$$

with a  $O(\cdot)$  uniform in  $k$ . Conclude that  $X_n$  converges in law to a standard Gaussian variable.

*Notice that  $X_n$  is a model for the number of disjoint cycles of a random permutation of  $n$  elements.*

## 2. MOD-CONVERGENCE AND BERRY-ESSEEN ESTIMATES

Lévy's criterion is mostly used in order to prove convergence towards a Gaussian random variable. In this second part we measure the difference between the distribution of the  $X_n$ 's and the Gaussian distribution under slightly stronger hypotheses than before. Hence, we consider a sequence of real-valued random variables  $(Y_n)_{n \in \mathbb{N}}$  such that

$$\forall n \in \mathbb{N}, \quad \forall t \in \mathbb{R}, \quad \mathbb{E}[e^{itY_n}] = \phi_n(t) = e^{-\frac{\lambda_n t^2}{2}} \psi_n(t); \quad (4)$$

$$\lambda_n \geq 0, \quad \lim_{n \rightarrow \infty} \lambda_n = +\infty; \quad (5)$$

$$\lim_{n \rightarrow \infty} \psi_n = \psi \quad \text{at speed } o\left(\frac{1}{\sqrt{\lambda_n}}\right). \quad (6)$$

In Equation (4), we call the functions  $\psi_n(t)$  the *residues* of the characteristic functions  $\phi_n(t)$ . We assume that they are continuously differentiable functions on the real line, and that  $\psi_n$  converge uniformly and sufficiently fast on every compact to the function  $\psi$ , which

is itself continuously differentiable (and with  $\psi_n(0) = \psi(0) = 1$ ); this is the meaning of Equation (6). So,

$$\forall \varepsilon > 0, \forall T > 0, \exists N, \forall n \geq N, \sup_{t \in [-T, T]} |\psi_n(t) - \psi(t)| \leq \frac{\varepsilon}{\sqrt{\lambda_n}}.$$

- (1) Set  $X_n = Y_n/\sqrt{\lambda_n}$ . Show that  $X_n$  converges in law to a standard Gaussian variable of mean 0 and variance 1. For this reason, Hypotheses (4)-(6) are called hypotheses of *mod-Gaussian convergence*.
- (2) For two probability measures  $\mu$  and  $\nu$  on  $\mathbb{R}$ , one defines their Kolmogorov distance as

$$d(\mu, \nu) = \sup_{x \in \mathbb{R}} |F_\mu(x) - F_\nu(x)|,$$

where  $F_\mu(x)$  is the cumulative distribution function of  $\mu$ , that is  $\mu(-\infty, x)$ . Prove that if  $\mu$  is absolutely continuous with respect to the Lebesgue measure, then

$$\mu_n \rightarrow \mu \iff d(\mu_n, \mu) \rightarrow 0.$$

One can use freely Dini's theorem, which says that bounded increasing functions that converge pointwise converge in fact uniformly (optional: prove Dini's theorem).

Hence, in the following, we shall measure the convergence  $X_n \rightarrow \mathcal{N}(0, 1)$  by computing  $d(\mu_{Y_n}, \mathcal{N}_{(0, \lambda_n)}) = d(\mu_{X_n}, \mathcal{N}_{(0, 1)})$ . For  $T > 0$ , we set

$$\Delta_T(x) = \frac{1 - \cos(Tx)}{\pi T x^2}.$$

Questions 3. to 5. are devoted to a proof of Berry's lemma, which relates the Kolmogorov distance to the behavior of characteristic functions. Then, in Questions 6. to 8., we apply this lemma to the situation of mod-convergence.

- (3) Fix  $T > 0$ . Show that  $\Delta_T(x) \geq 0$  for all  $x \in \mathbb{R}$ ; that  $\int_{\mathbb{R}} \Delta_T(x) dx = 1$ ; and that

$$\widehat{\Delta_T}(t) = \int_{\mathbb{R}} \Delta_T(x) e^{itx} dx = \begin{cases} 1 - \frac{|t|}{T} & \text{if } |t| \leq T, \\ 0 & \text{otherwise.} \end{cases}$$

Show also that the probability measure  $\Delta_T(x) dx$  gives to the set  $\{x, |x| \geq h\}$  a mass smaller than  $\frac{4}{\pi T h}$ .

- (4) Let  $F$  be the cumulative distribution function of a probability measure, and  $G$  be a bounded function with

$$\lim_{x \rightarrow -\infty} G(x) = 0 \quad ; \quad \lim_{x \rightarrow \infty} G(x) = 1 \quad ; \quad |G'(x)| \leq m$$

for a certain constant  $m$ . We denote

$$\begin{aligned} D(x) = F(x) - G(x) & \quad ; \quad D_T(x) = \int_{\mathbb{R}} D(x-y) \Delta_T(y) dy; \\ \eta = \sup_{x \in \mathbb{R}} |D(x)| & \quad ; \quad \eta_T = \sup_{x \in \mathbb{R}} |D_T(x)|. \end{aligned}$$

If  $\eta = 0$ , show that  $\eta_T = 0$ . Otherwise, we fix an element  $x_0$  such that  $|D(x_0)| = \eta$ ; for instance we assume  $D(x_0) = \eta$ . Show that if  $h = \frac{\eta}{2m}$  and  $x = x_0 + h$ , then

$$D(x-y) \geq \frac{\eta}{2} + my \quad \text{for all } |y| \leq h.$$

Prove then that

$$\eta_T \geq \frac{\eta}{2} - \frac{12m}{\pi T}$$

(distinguish the cases  $\eta = 0$  and  $\eta > 0$ , and in this case split the integral  $D_T(x)$  in two parts).

- (5) Suppose that  $F = F_\mu$  and  $G$  have characteristic functions

$$\phi_\mu(t) = \int_{\mathbb{R}} e^{itx} \mu(dx) \quad ; \quad \phi_G(t) = \int_{\mathbb{R}} e^{itx} g(x) dx = \widehat{g}(t) \quad \text{with } g = G'.$$

Using Fourier inversion formula  $f(x) = \frac{1}{2\pi} \int_{\mathbb{R}} \widehat{f}(t) e^{-itx} dt$ , prove that

$$\begin{aligned} D_T(x) &= \frac{1}{2\pi} \int_{\mathbb{R}} \widehat{\Delta}_T(t) \left( \frac{\phi_\mu(t) - \phi_G(t)}{-it} \right) e^{-itx} dt; \\ \eta &\leq \frac{1}{\pi} \int_{-T}^T \left| \frac{\phi_\mu(t) - \phi_G(t)}{t} \right| dt + \frac{24m}{\pi T}. \end{aligned} \tag{7}$$

This last inequality is *Berry's lemma*.

- (6) Under the hypotheses (4)-(6), show that

$$\phi_{\mu_n}(t) = \mathbb{E}[e^{itX_n}] = e^{-\frac{t^2}{2}} \left( 1 + \frac{\psi'(0)t}{\sqrt{\lambda_n}} + o\left(\frac{t}{\sqrt{\lambda_n}}\right) \right).$$

Prove also that  $e^{-\frac{t^2}{2}} \left( 1 + \frac{\psi'(0)t}{\sqrt{\lambda_n}} \right)$  is the Fourier transform of

$$g_n(x) = \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} \left( 1 - \frac{\psi'(0)ix}{\sqrt{\lambda_n}} \right).$$

We set  $G_n(x) = \int_{-\infty}^x g_n(y) dy$ ; check that this function satisfies the previous assumptions.

- (7) In the previous setting, prove that

$$\eta_n = \sup_{x \in \mathbb{R}} |F_{\mu_n}(x) - G_n(x)| = o\left(\frac{1}{\sqrt{\lambda_n}}\right)$$

(hint: take  $T = K\sqrt{\lambda_n}$  with  $K$  big, and split the integral in (7) in two parts

$$\begin{aligned} t &\in [-\varepsilon\sqrt{\lambda_n}, \varepsilon\sqrt{\lambda_n}]; \\ t &\in [-K\sqrt{\lambda_n}, K\sqrt{\lambda_n}] \setminus [-\varepsilon\sqrt{\lambda_n}, \varepsilon\sqrt{\lambda_n}]. \end{aligned}$$

with  $\varepsilon$  sufficiently small).

- (8) Assuming  $\psi'(0) \neq 0$ , show that

$$d(\mu_{X_n}, \mathcal{N}_{(0,1)}) = \frac{|\text{Im}(\psi'(0))| (1 + o(1))}{\sqrt{2\pi\lambda_n}}.$$

- (9) Optional: apply this result to the example of 1.(11).