# Forecasting with linear regression

Yannig Goude-yannig.goude@edf.fr

# **AGENDA**

- 1. FORECASTING: PRINCIPLES AND METHODOLOGY
- 2. LINEAR REGRESSION
- 3. APPLICATION TO ELECTRICITY LOAD FORECASTING

# **FORECASTING:** principles and methodology



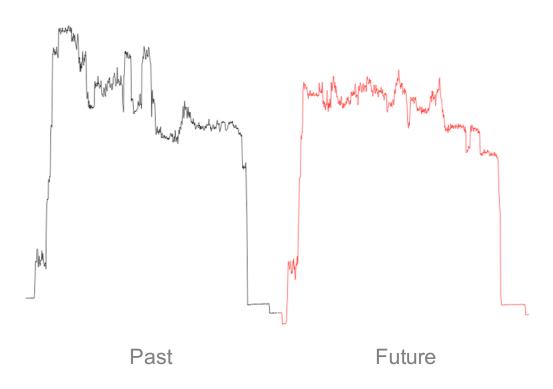
### **FORECASTING**

### Can take many forms:

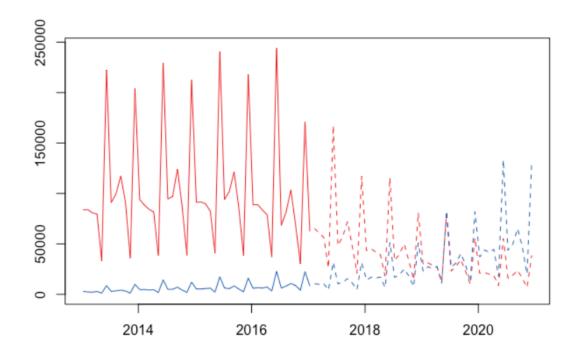
- Expert advice
- Prospective (science fiction, anticipation)
- Scenario generation
- What if scenarios
- Physical modeling

We will focus here on forecasting from data

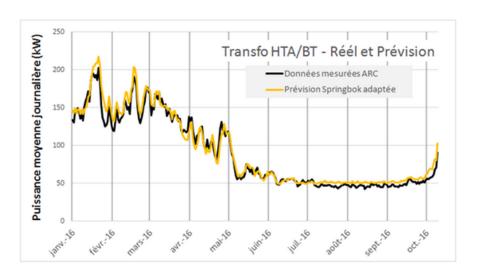
Forecasting electricity load (industrial consumption at 5 min resolution) at a one day horizon:



Forecasting car sales (UK) at a 2 years horizon:



Forecasting low voltage stations load:

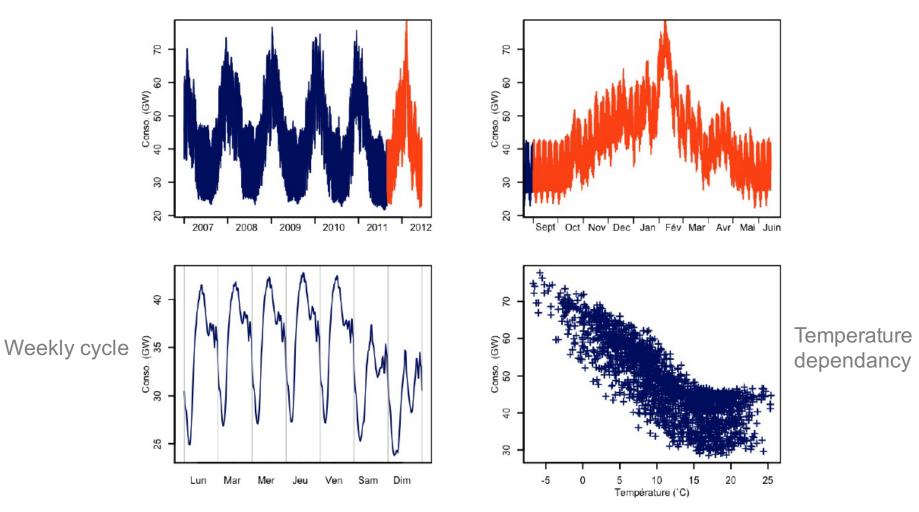


- Characterize statistical properties of the data:
  - Endogenous (time dependancy)
  - Exogenous (dependancy with other covariates)
- Correlation/causality
- Stationnarity
- Mean forecast/ distribution forecast/ quantile forecast

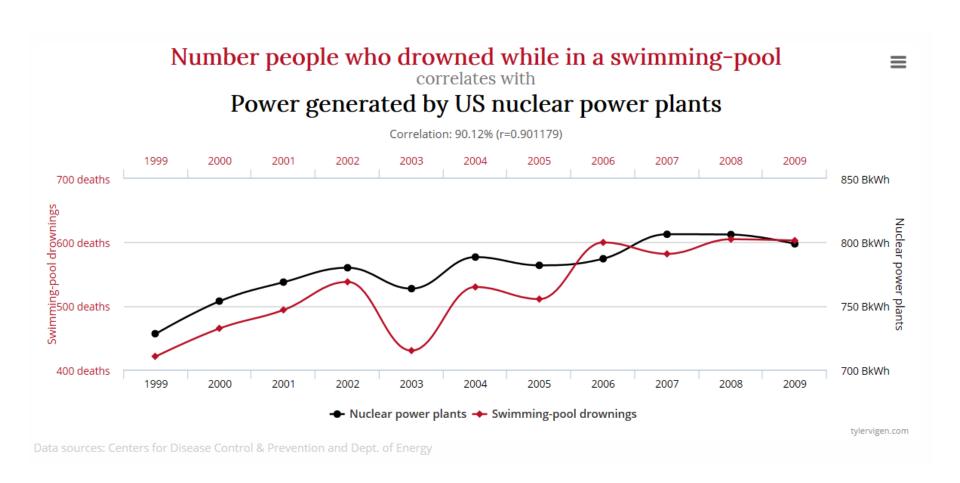
We restrict here to **statistical models** which are a good compromise between quality of the forecast and interpretability of the models.

Some recent development have been done in the field of **Al** methods (e.g. deep learning) but these approach are still **black** boxes.

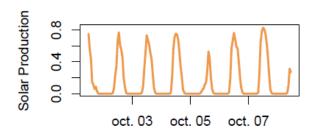
Exogenous/endogenous dependance: the exemple of french electricity load

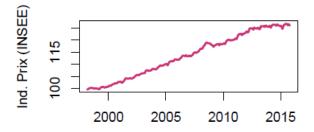


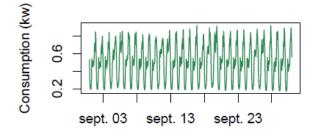
### Correlation/causality

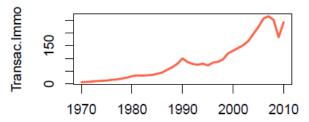


Stationnarity: is the law of the process stable with time?

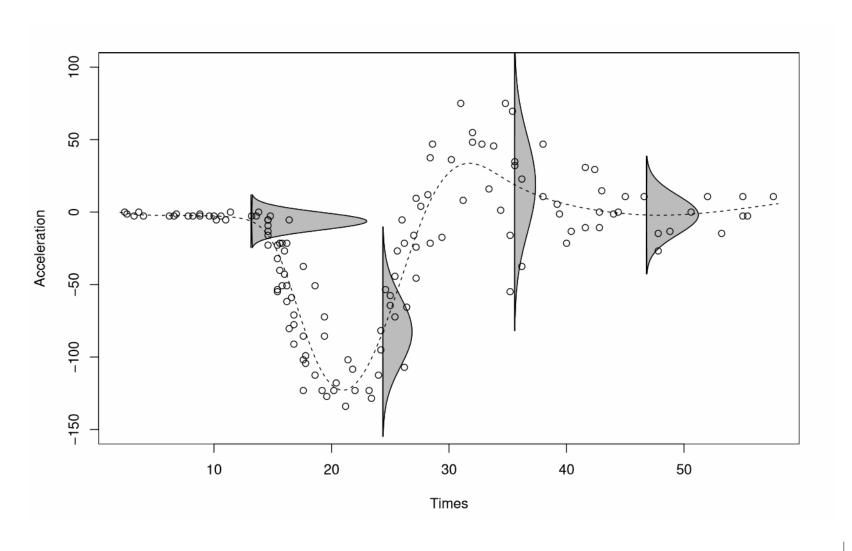








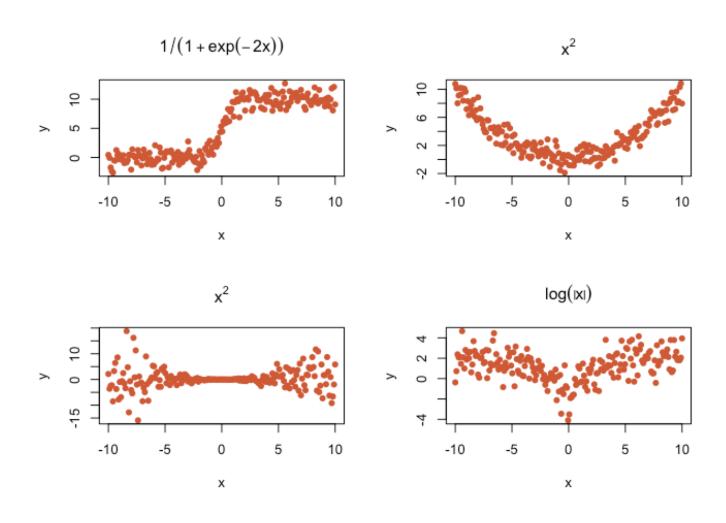
Mean forecast, density forecast, quantile forecast



# **FORECASTING:** descriptive statistics

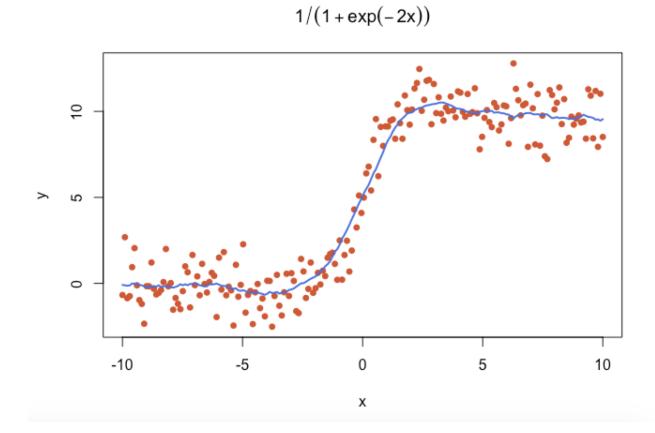


### Scatter plots

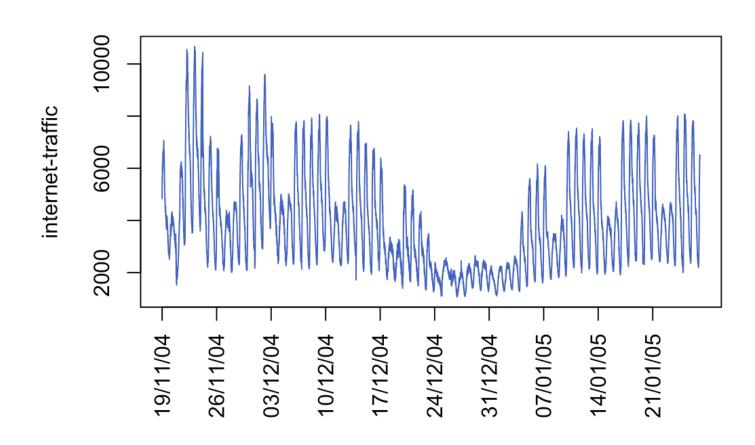


### Scatter plots / kernel smoothing

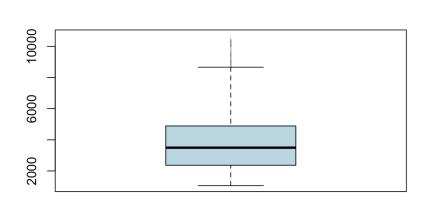
$$\widehat{f}_h(x) = \frac{\sum_{t=1}^n y_t K(\frac{x-t}{h})}{\sum_{t=1}^n K(\frac{x-t}{h})}$$

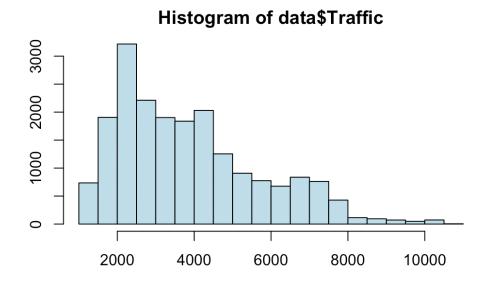


...or time plots



### Histograms/ boxplots





#### Autocorrelation

More specifically for time series, these statistics are useful:

- the **empirical mean** of a time series  $(y_t)_{1 \le t \le n}$ :  $\bar{y}_n = \frac{1}{n} \sum_{t=1}^n y_t$
- the empirical standart deviation to estimate its **dispersion** :  $\sigma_n = \sqrt{\frac{1}{n} \sum_{t=1}^n (y_t \bar{y}_n)^2}$
- empirical auto-covariance  $\gamma$  or autocorrelation  $\rho$  indicate the temporal (linear) **dependancy**:

$$\gamma_n(h) = \frac{1}{n} \sum_{t=1}^{n-h} (y_t - \bar{y}_n)(y_{t+h} - \bar{y}_n)$$

$$\rho_n(h) = \frac{\gamma_n(h)}{\gamma(0)}$$

remark that  $\gamma_n(0) = \sigma_n^2$ .

thus  $\rho_n(h)$  is the estimate of the correlation between  $y_t$  and  $y_{t+h}$ , supposing this correlation exists and is stable in function of time (stationnarity)

#### Partial Autocorrelation

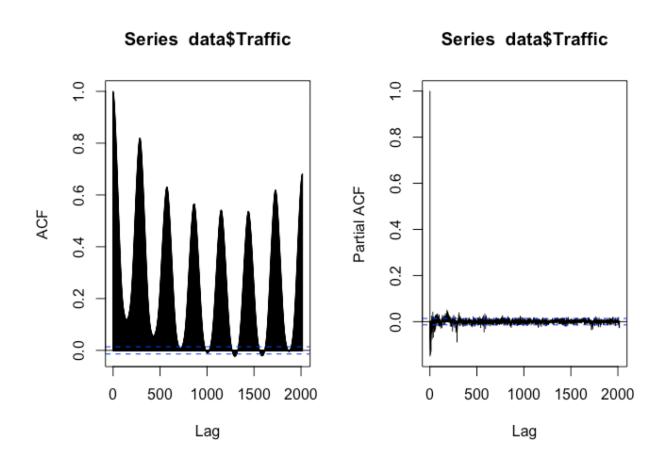
• partial autocorrelation (PACF): dependancy between two instant t and t+k conditionnally to what happened at times  $t+1,\ldots,t+k-1$ 

To obtain order h PACF one have to consider the following linear model:

$$y_t = \alpha_1 y_{t-1} + \alpha_2 y_{t-2} + \ldots + \alpha_h y_{t-h} + \varepsilon_t$$

the order h PACF is defined as  $\alpha_h$  and can be estimated solving the OLS problem.

#### ACF and PACF



## **FORECASTING:** linear model



We consider a target Y which is a real random variable that we want to forecast according to:

- paste value of *Y*
- and/or covariates  $X_1$ , ...,  $X_p$

We assume that the data are generated according to the following (linear regression) model:

$$y_i = x_{i,1}\beta_1 + x_{i,2}\beta_2 + \ldots + x_{i,p}\beta_p + \varepsilon_i$$

for  $i=1,\ldots,n$  observations. Our objective is to estimate the unknown parameters  $\beta=(\beta_1,\ldots,\beta_p)$ 

We will also suppose:

- $\varepsilon_i$  are independent and indentically distributed (iid)
- mean 0 and constant variance:  $E(\varepsilon_i) = 0$ ,  $V(\varepsilon_i) = \sigma^2$

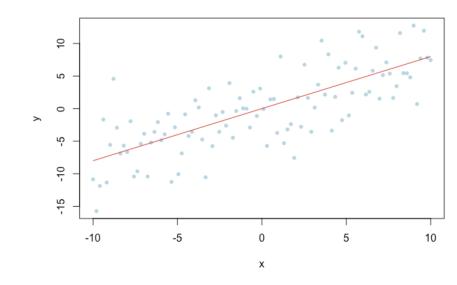
This can be rewritten with matrix notations:

$$Y = X\beta + \varepsilon$$

where



- $Y \in \mathcal{R}^p$ ,  $\beta \in \mathcal{R}^p$
- $E(\varepsilon) = 0$ ,  $V(\varepsilon_i) = \sigma^2$



suppose that we measure our performance with the quadratic loss, we can solve the well known ordinary least square problem to infer  $\beta$  from the data:

$$\min_{\beta \in \mathcal{R}^p} \sum_{i=1}^n (y_i - x_i \beta)^2 = \min_{\beta \in \mathcal{R}^p} ||Y - X\beta||^2$$

This can be seen as a convex optimisation problem where we minimise the fonction  $g: \beta \to \sum (y_i - x_i \beta)^2$ .

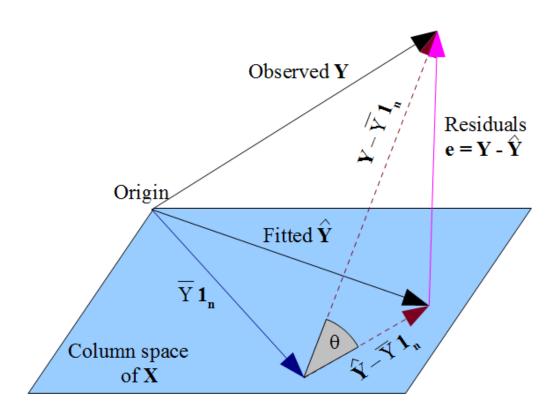
So,  $\widehat{\beta}$  satisfying  $\frac{dg}{d\beta}(\widehat{\beta})=0$ , ie:

$$\frac{dg}{d\beta} = -2X^T Y + 2X^T X \beta = 2X^T (X\beta - Y)$$

thus,  $X\widehat{\beta} = Y$  and  $X^T X\widehat{\beta} = X^T Y$  and, as X is rank p:

$$\widehat{\beta} = (X^T X)^{-1} X^T Y$$

### Geometric interpretation



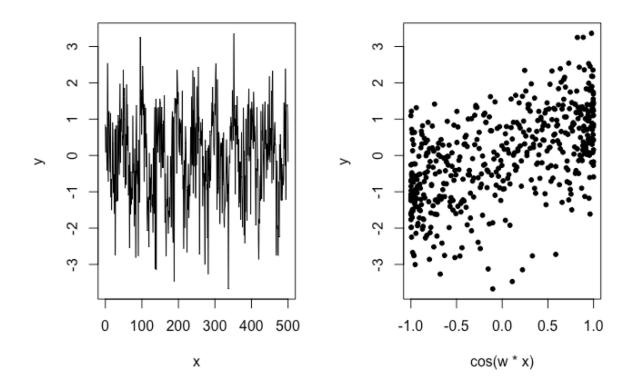
#### **Useful statistics:**

- $R^2 = \cos^2(\theta) = \frac{\|\widehat{Y} \overline{Y}\|^2}{\|Y \overline{Y}\|^2} = 1 \frac{\|Y \widehat{Y}\|^2}{\|Y \overline{Y}\|^2}$  the proportion of variance explained by our model. It indicates wether the regression is close to the observations (including the noise variance).
- don't work if we compare different nature of models (e.g. multiplicative vs additive)
- can induce overfitting as  $R^2$  increases as p increases, there is an adjusted version to take dimension p into account:  $R^2_a = 1 \frac{n}{n-p} \frac{||Y \widehat{Y}||^2}{||Y \overline{Y}||^2}$

Linear models are a powerful tool as, conditional to the good transformations of the data  $X \to Xnew$ , we can often express Y as a linear combination of Xnew. Here are a few examples.

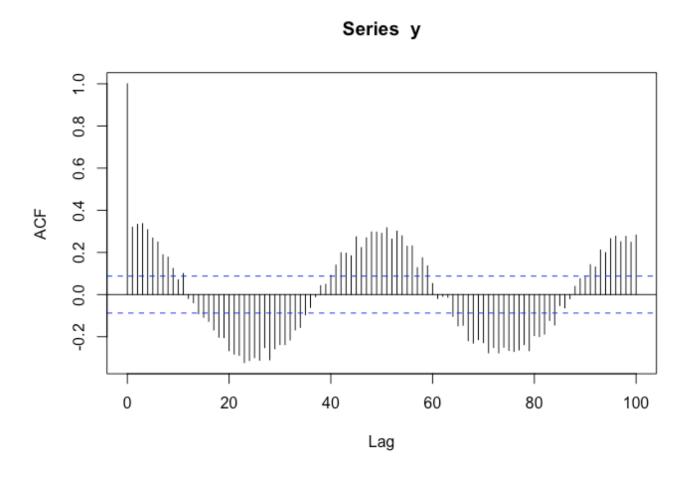
• periodic data: Fourier basis regression

For a well chosen  $\omega = 2 * \pi/T$  and an harmonic  $k: Xnew = \cos(k * \omega x)$ 

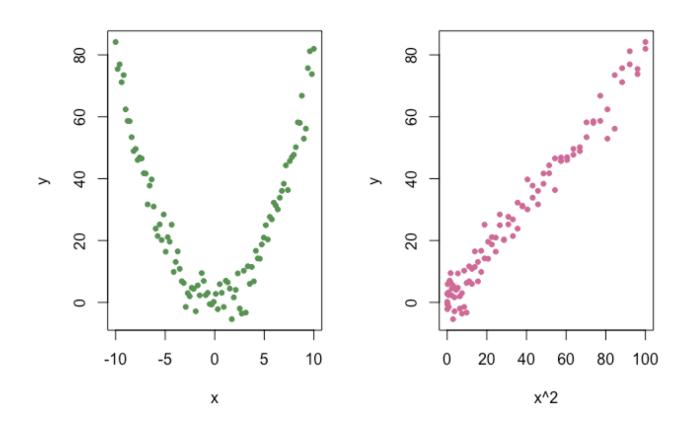


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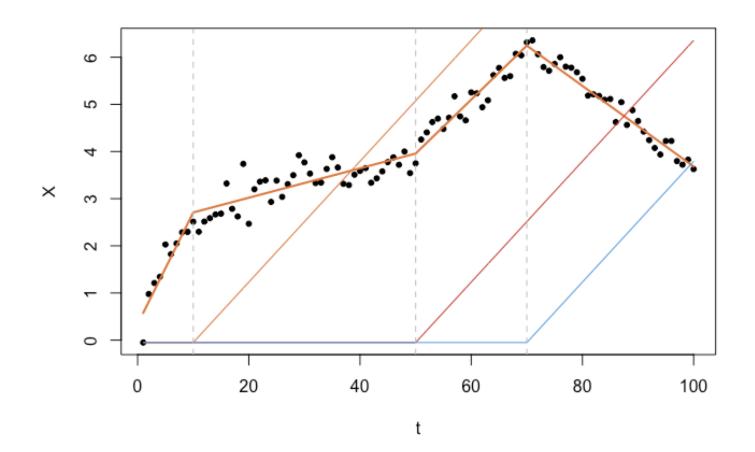
 $\boldsymbol{\omega}$  can be chosen according to a frequency analysis of the signal:



• polynomial transformation:



• spline basis decomposition: e.g. truncated power functions  $1, x, (x - k)_+$ 



The forecaster needs an objective criteria to:

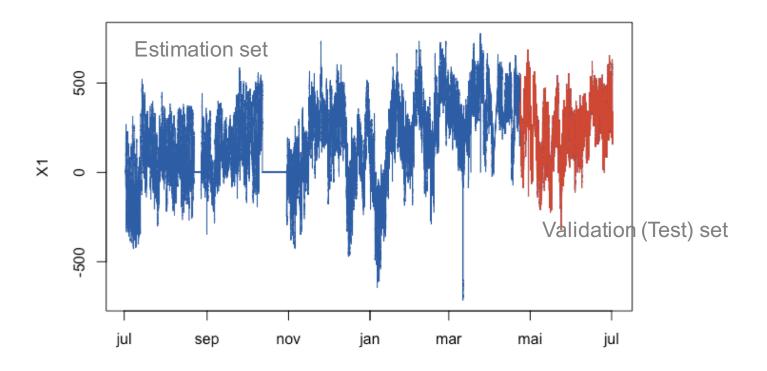
- Select the set of covariates to include into the model
- Find the good transformation of the covariates
- Calibrate the model
- Have a good estimate of its forecasting performances



#### Many pitfalls:

- Overfitting
- Extrapolation problem (trends)
- Most of the time data are not iid

#### Test set



### Work only if:

- the data have the same generation process in the 2 sets
- we have enough data to split it

Try to choose the test set in accordance with your final purpose/ the characteristic of the data

#### Cross validation

#### leave one out:

- choose randomly  $i \in \{1, \ldots, n\}$
- fit your model on all the data except i ie  $(x_1, y_1), \ldots, (x_{i-1}, y_{i-1}), (x_{i+1}, y_{i+1}), \ldots, (x_n, y_n)$ , denotes this model  $\widehat{\phi}_{-i}$
- estimate a forecast error  $(y_i \widehat{\phi}_{-i}(X_i))^2$
- ullet repeat that N times and compute an estimate of the forecast error of your model  $\phi$

$$R_N(\phi) = \frac{1}{N} \sum_{i=1}^{N} (y_i - \widehat{\phi}_{-i}(X_i))^2$$

#### Cross validation

K-fold:

For  $k \in \{1, ..., K\}$ ,

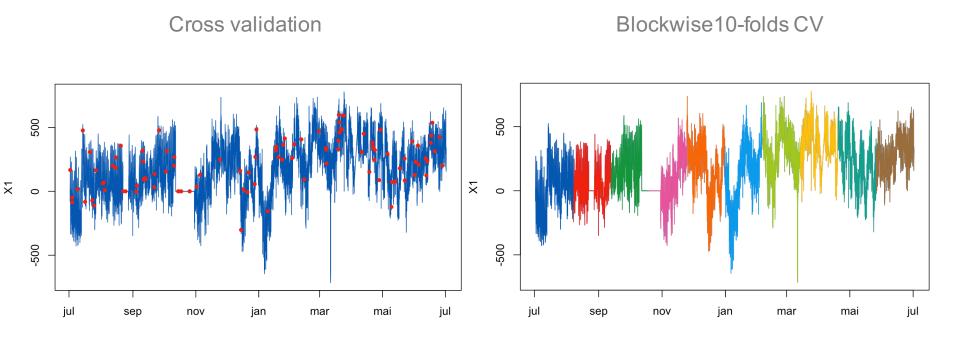
- choose randomly  $I_k = (i_1, \dots, i_Q) \in \{1, \dots, n\}^Q$ , where K \* Q = n
- ullet fit your model on all the data except  $I_k$  , denotes this model  $\widehat{\phi}_{-I_k}$
- estimate a forecast error  $R_{I_k} = \frac{1}{Q} \sum_{k=1}^{Q} (y_{i_k} \widehat{\phi}_{-I_k}(x_{i_k}))^2$

then compute an estimate of the forecast error of your model  $\phi$ 

$$R_K(\phi) = \frac{1}{K} \sum_{i=1}^K R_{I_k}$$

#### Remarks:

- $I_k$  is here randomly sample but it could be blocks of consecutives observations (blockwise K-fold CV) so that  $I_1,...,I_K$  is a partition of 1,...,n. This is particularly relevant in the time series context.
- in addition to the error  $R_K(\phi)$ , if K is sufficiently large, we can also compute a measure of the uncertainty of this estimate (variance, quantile...)



Sequential testing

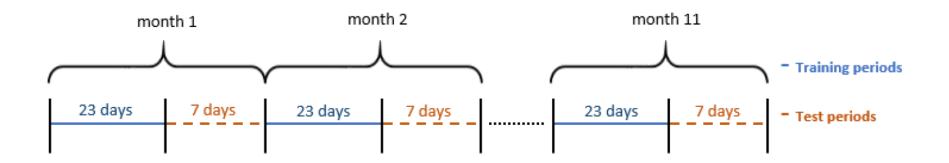
For  $t \in \{n_0, ..., n\}$ :

- fit a model  $\widehat{\phi}_t$  on the data  $(x_1, y_1), \dots, (x_t, y_t)$
- forecast  $y_{t+1}$  as  $\widehat{\phi}_t(x_{t+1})$

then compute an estimate of the forecast error of your model  $\phi$ 

$$R_{n_0}(\phi) = \frac{1}{n - n_0} \sum_{t=n_0+1}^{n} (y_t - \widehat{\phi}_t(x_t))^2$$

Well chosen blockwise test set



Yearly seasonnal time series

For linear model we have this convenient property:

$$CV = \frac{1}{n} \sum_{i=1}^{n} \frac{(y_i - \widehat{y}_i)^2}{(1 - H_{i,i})^2} \qquad \widehat{\varepsilon}_i^{-i} = \widehat{\varepsilon}_i / (1 - H_{i,i})$$

#### Preuve:

lemme d'inversion matriciel: Soit M une matrice symétrique inversible  $p \times p$ , u et v deux vecteurs de taille p. Alors:

$$(M + uv')^{-1} = M^{-1} - \frac{M^{-1}uv'M^{-1}}{1 + u'M^{-1}v}$$

- $X'X = X'_{-i}X_{-i} + x_ix_i'$
- $X'Y = X'_{-i}Y_{-i} + x_iy_i$
- $h_{i,i} = x_i'(X'X)^{-1}x_i$

#### Résiduals checks

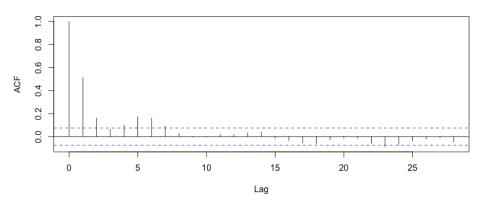
- Independance: acf, pacf
- Test de Box-Pierce

$$H_0(h): \rho_{\varepsilon}(1) = \rho_{\varepsilon}(2) = \dots = \rho_{\varepsilon}(h) = 0$$

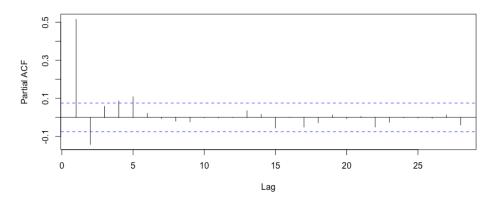
$$H_1(h): \exists k \in (1,...,h)$$
t.  
q $\rho_\varepsilon(k) \neq 0$ 

$$Q_{BP}(h) = n \sum_{j=1}^{h} \widehat{\rho}_{\varepsilon}(j)^2 \xrightarrow[n \to +\infty]{\mathcal{L}} \chi^2(h-k)$$

#### Series eps

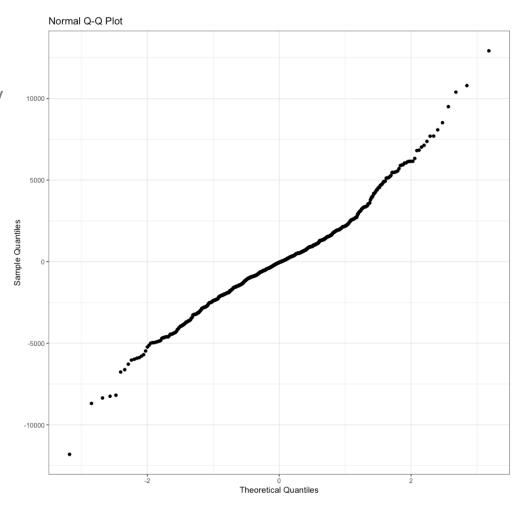


#### Series eps

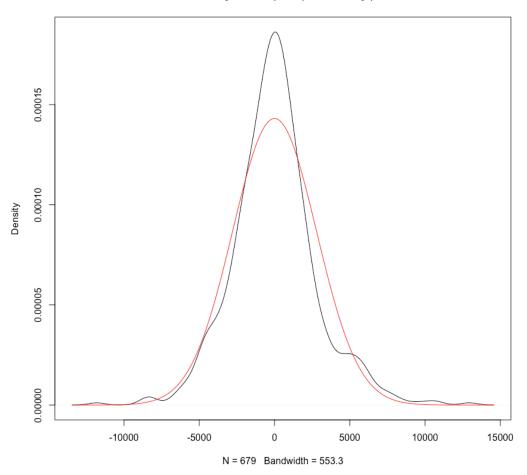


### Adequation to a given distribution:

- Qqplot
- Density estimation
- Tests: chi2, kolmogorov-Smirnov



#### density.default(x = eps, bw = "sj")



Now, linear modeling of french electricity consumption....